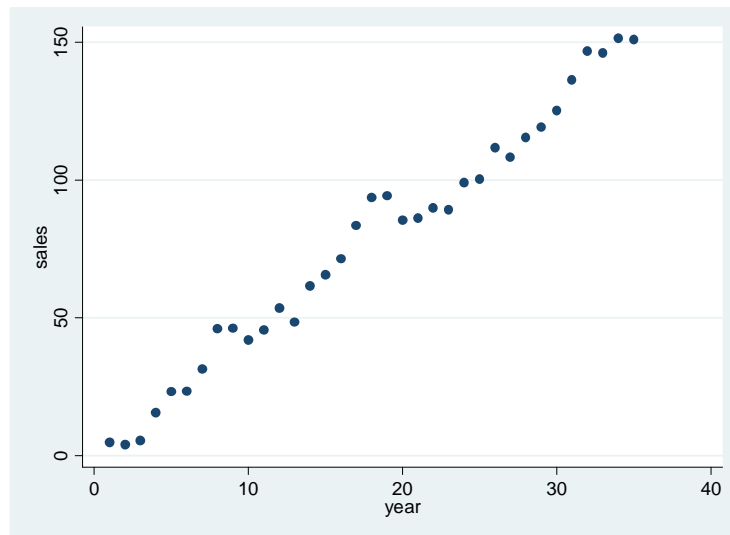


MA397 – Regression-Based Methods for Time Series

Moving Average Smoother

For this exercise we will be using the *sales.dta* dataset found on the course webpage at <http://www.colby.edu/personal/l/lobrien/ma397.html>.

Recall that the data consist of sales information over a 35-year period. The plot of the data are below:



We see a clear seasonal trend. There may be a seasonal component as well, but it is less clear. We first fit the model using OLS and calculate the Durbin-Watson statistic to see if there is residual autocorrelation (make sure to set the data as time series first).

```
. tsset year
      time variable:  year, 1 to 35
             delta:  1 unit

. regress sales year

-----+-----
Source |           SS       df       MS              Number of obs =      35
-----+-----+-----+-----
Model |  65875.2068         1  65875.2068          F( 1,   33) = 1615.72
Residual |  1345.45362        33   40.7713217          Prob > F      =  0.0000
-----+-----+-----+-----
Total |  67220.6604        34  1977.07825          R-squared     =  0.9800
                                           Adj R-squared =  0.9794
                                           Root MSE    =  6.3852

-----+-----
sales |           Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----+-----+-----
year |    4.29563     .1068669    40.20   0.000     4.078208    4.513053
_cons |    .4015129    2.205708     0.18   0.857    -4.086034    4.88906

. estat dwatson

Durbin-Watson d-statistic( 2,   35) = .8207266
```

We see that the Durbin-Watson statistic is 0.8207266, which is less than the lower cutoff of 1.40 indicating positive autocorrelation. So let's fit a first-order autoregressive model to the data.

To fit autoregressive, moving average, or autoregressive-moving average models in Stata go to **Statistics > Times series > ARIMA and ARMAX models**. A wide variety of models can be fit via this option. For a simple first-order AR model, enter the dependent variable (sales) and the independent variables (year) in the appropriate boxes. Enter the autoregressive order of 1 (leave the difference and moving average orders at 0. Click okay to obtain:

```
. arima sales year, arima(1,0,0)

(setting optimization to BHHH)
Iteration 0: log likelihood = -106.26769
Iteration 1: log likelihood = -106.2573
Iteration 2: log likelihood = -106.25527
Iteration 3: log likelihood = -106.25519
Iteration 4: log likelihood = -106.25501
(swapping optimization to BFGS)
Iteration 5: log likelihood = -106.25321
Iteration 6: log likelihood = -106.25321

ARIMA regression

Sample: 1 - 35                                Number of obs   =          35
Log likelihood = -106.2532                    Wald chi2(2)    =       346.62
                                              Prob > chi2     =         0.0000
```

		Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
sales							
year		4.295973	.2378623	18.06	0.000	3.829772	4.762175
_cons		.4060005	4.988835	0.08	0.935	-9.371937	10.18394
ARMA							
ar							
L1.		.5735942	.1496151	3.83	0.000	.2803541	.8668344
/sigma		5.008673	.7330814	6.83	0.000	3.57186	6.445486

Note the difference in the standard errors of the regression coefficients (they were underestimated with OLS), but that the OLS estimates were unbiased. The coefficient for the ARMA output is phi (0.574 in this case). The “sigma” value is the root MSE for the model. The other statistics given in the textbook printout are not available since the model was estimated using maximum likelihood rather than the modified OLS procedure.

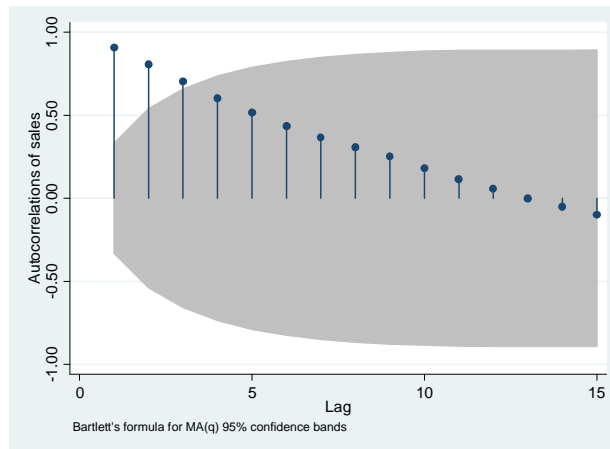
Stata has many interesting commands for time series data. For example, to calculate the autocorrelations using the command:

```
. corrgram sales
```

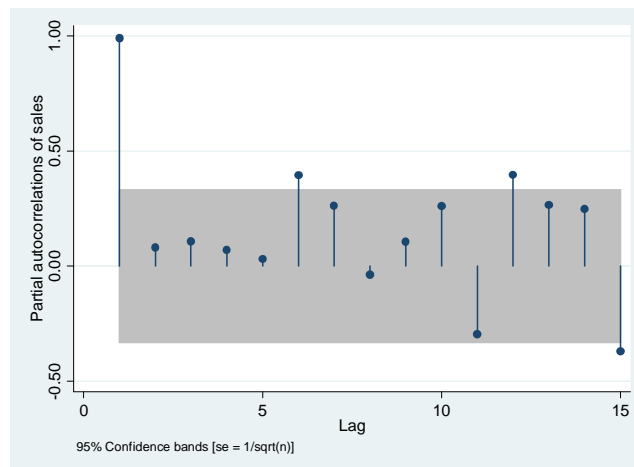
LAG	AC	PAC	Q	Prob>Q	-1	0	1	-1	0	1
					[Autocorrelation]			[Partial Autocor]		
1	0.9077	0.9914	31.384	0.0000						
2	0.8058	0.0809	56.867	0.0000						
3	0.7036	0.1081	76.901	0.0000						
4	0.6032	0.0707	92.103	0.0000						
5	0.5143	0.0313	103.52	0.0000						
6	0.4351	0.3940	111.97	0.0000						

7	0.3656	0.2616	118.16	0.0000	--	--
8	0.3072	-0.0359	122.68	0.0000	--	--
9	0.2508	0.1065	125.82	0.0000	--	--
10	0.1805	0.2607	127.5	0.0000		--
11	0.1152	-0.2948	128.22	0.0000		--
12	0.0566	0.3962	128.4	0.0000		--
13	-0.0011	0.2653	128.4	0.0000		--
14	-0.0504	0.2493	128.56	0.0000		
15	-0.1000	-0.3711	129.21	0.0000		--

We can see this graphically with a correlogram by giving the command: `ac sales`



A partial autocorrelation plot can be obtained by giving the command: `pac sales`



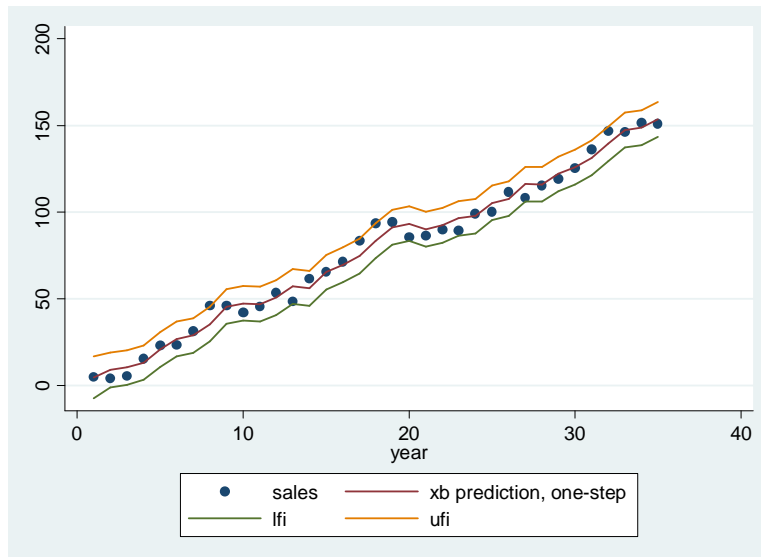
From these plots it looks as if our first-order AR model is a good fit. We may now want to forecast using it. Forecasting in Stata is straightforward after fitting a time series model. However, forecasting beyond the end of the x-data range is not possible. To get the point estimates simply use the “predict” command.

```
. predict yhat
(option xb assumed; predicted values)
```

If we want to generate confidence intervals around these, we need to first estimate the MSE of the observations.

```
. predict mse, mse  
. gen lfi = yhat-2*sqrt(mse)  
. gen ufi=yhat+2*sqrt(mse)
```

The variables *lfi* and *ufi* now contain the confidence bands around the forecasted values.



We can try to estimate seasonal variation noticing that there seems to be a 10-year cycle in the data. Try to fit sine and cosine functions and rerun the model:

```
. gen sine=sin((_pi/10)*year)  
. gen cosine=cos((_pi/10)*year)  
. arima sales year sine cosine, arima(1,0,0)  
  
(setting optimization to BHHH)  
Iteration 0: log likelihood = -104.43414  
Iteration 1: log likelihood = -104.42012  
Iteration 2: log likelihood = -104.4149  
Iteration 3: log likelihood = -104.41449  
Iteration 4: log likelihood = -104.41442  
(switching optimization to BFGS)  
Iteration 5: log likelihood = -104.41442
```

ARIMA regression

```
Sample: 1 - 35                               Number of obs   =      35
Log likelihood = -104.4144                    Wald chi2(4)    =    581.77
                                              Prob > chi2     =    0.0000
```

sales	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
sales						
year	4.195335	.1850473	22.67	0.000	3.832649	4.558021
sine	-4.362489	2.234392	-1.95	0.051	-8.741816	.0168381
cosine	.1228667	1.974232	0.06	0.950	-3.746556	3.992289
_cons	2.468488	3.579696	0.69	0.490	-4.547589	9.484564
ARMA						
ar						
L1.	.4631711	.189805	2.44	0.015	.0911601	.8351822
/sigma	4.763073	.6386057	7.46	0.000	3.511429	6.014717

```
. arima sales year sine, arima(1,0,0)
```

```
(setting optimization to BHHH)
Iteration 0: log likelihood = -104.43257
Iteration 1: log likelihood = -104.4219
Iteration 2: log likelihood = -104.41995
Iteration 3: log likelihood = -104.41949
Iteration 4: log likelihood = -104.41902
(switching optimization to BFGS)
Iteration 5: log likelihood = -104.4162
Iteration 6: log likelihood = -104.41619
Iteration 7: log likelihood = -104.41618
```

ARIMA regression

```
Sample: 1 - 35                               Number of obs   =      35
Log likelihood = -104.4162                    Wald chi2(3)    =    575.80
                                              Prob > chi2     =    0.0000
```

sales	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
sales						
year	4.193308	.1850032	22.67	0.000	3.830708	4.555907
sine	-4.363483	2.224331	-1.96	0.050	-8.723091	-.0038739
_cons	2.493923	3.523166	0.71	0.479	-4.411356	9.399202
ARMA						
ar						
L1.	.463647	.1857311	2.50	0.013	.0996207	.8276732
/sigma	4.763267	.636693	7.48	0.000	3.515372	6.011163

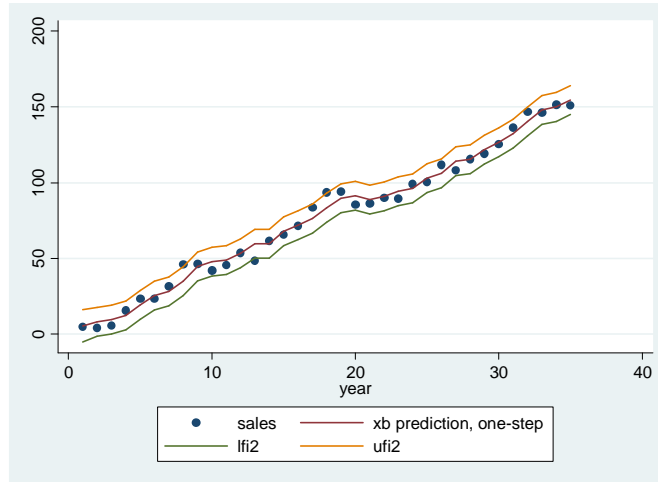
We see that the sine function may provide better forecasts. We can compare the forecasts by generating new values and examining the plot,

```
. predict yhat2
(option xb assumed; predicted values)

. predict mse2, mse

. gen lfi2 = yhat2-2*sqrt(mse2)

. gen ufi2=yhat2+2*sqrt(mse2)
```



The smaller root MSE for the seasonally-adjusted model indicates a slightly better fit, but the difference appear to be small.