


## What if We Don't Know $\boldsymbol{\sigma}$ ?

- We usually don't know the population standard deviation, $\sigma$.
- Estimate $\sigma$ using the sample standard deviation s.
- This will change our formula for the confidence interval since we will no longer be under a normal distribution (the CLT doesn't help us here).


## Confidence Intervals ( $\sigma$ unknown)

We can estimate the sd of $\bar{x}$ using $\frac{s}{\sqrt{n}}$; this is called the standard error of $\bar{x}$.
Estimating $\sigma$ with $s$ introduces a new source of variation. As a result the CI needs to be wider. How much wider?

## Confidence Intervals ( $\sigma$ unknown)

- When n is small, s doesn't estimate $\sigma$ well, so the CI needs to be quite a bit wider.
- When n is large, s estimates $\boldsymbol{\sigma}$ better, and the CI only needs to be slightly wider.
- How do we make the interval wider?
- Make the multiplier $z^{*}$ larger; use the $t-$ distribution ( $\mathrm{t}^{*}$ ) instead of the standard normal distribution.


## Student's t-distribution

- Properties of the t-distribution
- "Bell-shaped" and symmetric similar to the normal distribution.
- More spread out than the normal distribution.
- Exact shape depends on its degrees of freedom.
- As the number if degrees of freedom increases, the corresponding t-distribution looks more like a standard normal distribution.

Normal versus t-distribution


## Confidence Intervals ( $\sigma$ unknown)

Before $z=\frac{\bar{x}-\mu}{\sigma / \sqrt{n}}$ had a standard normal
distribution (CLT).
Now $t=\frac{\bar{x}-\mu}{s / \sqrt{n}}$ has a $t$-distribution with
$n-1$ degrees of freedom (df)
$\mathrm{df}=$ amount of information available in
data for estimating $\sigma$.

## Sampling Distribution of a Sample Mean

- Distribution of values taken by the sample mean in all possible samples of size $\mathbf{n}$ from the population with unknown $\sigma$.
population with unknown $\sigma$.
- For sample 1: SRS of size $\mathrm{n} \rightarrow \bar{x}_{1}, s_{1}, \frac{\bar{x}_{1}-\mu}{s_{1} / \sqrt{n}}$
- For sample 2: SRS of size $\mathrm{n} \rightarrow \bar{x}_{2}, s_{2}, \frac{\bar{x}_{2}-\mu}{s_{2} / \sqrt{n}}$
- Etc.


## CI's from the t-distribution

- Before, we had a CI for $\mu$ given by, $\bar{x} \pm z^{*} \frac{\sigma}{\sqrt{n}}$
Now the CI is given by
$\bar{x} \pm t^{*} \frac{s}{\sqrt{n}}$ and $t^{*} \frac{s}{\sqrt{n}}$ is the margin of error.
Note: This interval is exact when the underlying population has a normal distribution, but is approx. correct when n is "large."


## Example: Housing Prices

- In an SRS of size 25 , we obtain a sample mean of 215 , and a sample sd of of 42 .
$95 \% \mathrm{CI}: \bar{x} \pm t^{*} \frac{s}{\sqrt{n}}$
$=215 \pm 2.064 \frac{42}{\sqrt{25}}$
$=(197.7,232.4)$
Note with $\mathrm{n}=25, \mathrm{df}=\mathrm{n}-1$, and $t^{*}=2.064$


| 19 | 0.688 | 0.861 | 1.066 | 1.328 | 1.729 | 2.093 | 2.205 | 2.539 | 2.861 | 3.174 | 3.519 | 3.883 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 20 | 0.687 | 0.860 | 1.064 | 1.325 | 1.725 | 2.086 | 2.197 | 2.528 | 2.845 | 3.153 | 3.552 | 3.850 |
| 21 | 0.686 | 0.859 | 1.063 | 1.323 | 1.721 | 2.080 | 2.189 | 2.518 | 2.831 | 3.135 | 3.527 | 3.819 |
| 22 | 0.686 | 0.858 | 1.061 | 1.321 | 1.717 | 2.074 | 2.183 | 2.508 | 2.819 | 3.119 | 3.505 | 3.792 |
| 23 | 0.685 | 0.858 | 1.060 | 1.319 | 1.714 | 2.069 | 2.177 | 2.500 | 2.807 | 3.104 | 3.485 | 3.768 |
| 24 | 0.685 | 0.857 | 1.059 | 1.318 | 1.711 | 2.064 | 2.172 | 2.492 | 2.797 | 3.091 | 3.467 | 3.745 |
| 25 | 0.684 | 0.856 | 1.058 | 1.316 | 1.708 | 2.060 | 2.167 | 2.485 | 2.787 | 3.078 | 3.450 | 3.725 |
| 26 | 0.684 | 0.856 | 1.058 | 1.315 | 1.706 | 2.056 | 2.162 | 2.479 | 2.779 | 3.067 | 3.435 | 3.707 |
| 27 | 0.684 | 0.855 | 1.057 | 1.314 | 1.703 | 2.052 | 2.158 | 2.473 | 2.771 | 3.057 | 3.421 | 3.690 |
| 28 | 0.683 | 0.855 | 1.056 | 1.313 | 1.701 | 2.048 | 2.154 | 2.467 | 2.763 | 3.047 | 3.408 | 3.674 |
| 29 | 0.683 | 0.854 | 1.055 | 1.311 | 1.699 | 2.045 | 2.150 | 2.462 | 2.756 | 3.038 | 3.396 | 3.659 |
| 30 | 0.683 | 0.854 | 1.055 | 1.310 | 1.697 | 2.042 | 2.147 | 2.457 | 2.750 | 3.030 | 3.385 | 3.646 |
| 40 | 0.681 | 0.851 | 1.050 | 1.303 | 1.684 | 2.021 | 2.123 | 2.423 | 2.704 | 2.971 | 3.307 | 3.551 |
| 50 | 0.679 | 0.849 | 1.047 | 1.299 | 1.676 | 2.009 | 2.109 | 2.403 | 2.678 | 2.937 | 3.261 | 3.496 |
| 60 | 0.679 | 0.848 | 1.045 | 1.296 | 1.671 | 2.000 | 2.099 | 2.390 | 2.660 | 2.915 | 3.232 | 3.460 |
| 80 | 0.678 | 0.846 | 1.043 | 1.292 | 1.664 | 1.990 | 2.088 | 2.374 | 2.639 | 2.887 | 3.195 | 3.416 |
| 100 | 0.677 | 0.845 | 1.042 | 1.290 | 1.660 | 1.984 | 2.081 | 2.364 | 2.626 | 2.871 | 3.174 | 3.390 |
| 1000 | 0.675 | 0.842 | 1.037 | 1.282 | 1.646 | 1.962 | 2.056 | 2.330 | 2.581 | 2.813 | 3.098 | 3.300 |
| $z^{*}$ | 0.674 | 0.841 | 1.036 | 1.282 | 1.645 | 1.960 | 2.054 | 2.326 | 2.576 | 2.807 | 3.091 | 3.291 |
|  | 50\% | 60\% | 70\% | 80\% | 90\% | 95\% | 96\% | 98\% | 99\% | 99.5\% | 99.8\% | 99.9\% |
|  | Confidence level C |  |  |  |  |  |  |  |  |  |  |  |

Note: For a t-distribution with $24 \mathrm{df}, 95 \%$ of the area falls between -2.064 and +2.064 .

## Finding t Multipliers in Stata

- Use the command line for this.
- To find the value that cuts off a certain area (p) to the right under a t-distribution with df degrees of freedom:
- display invttail(df, p)
- For example, "display ttail(24,.025)" gives 2.064.


## Sampling Distribution for a Proportion

- Recall: when n is large ( $\mathrm{np} \geq 10$ and $\mathrm{n}(1-\mathrm{p}) \geq$ 10) then
the sampling distribution of $\hat{p}=\frac{x}{n}$
is approximately normal with
$\mu_{\hat{p}}=p$
$\sigma_{\hat{p}}=\sqrt{\frac{p(1-p)}{n}}$



## CI for a Proportion

- A "traditional" $95 \%$ CI for p : $\hat{p} \pm 1.96 \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$
where $\hat{p}=\frac{x}{n}$ is calculated from our sample. The multiplier can of course be changed for any confidence level.


## CI for a Proportion

- There is a problem with this formulation when p is close to 0 or 1 .
- Adjustment: Pretend we have 4 additional observations, 2 successes and 2 failures.
$\tilde{p} \pm z^{*} \sqrt{\frac{\tilde{p}(1-\tilde{p})}{n+4}}$
where $\tilde{p}=\frac{x+2}{n+4}$ is the adjusted sample proportion.


## Example

- In an SRS of 100 adults, obtain $x=95$ who want better health insurance.
95\% CI for $\mathrm{p}: \tilde{p} \pm 1.96 \sqrt{\frac{\tilde{p}(1-\tilde{p})}{n+4}}$
$\tilde{p}=\frac{x+2}{n+4}=\frac{95+2}{100+4}=0.933$
$95 \%$ CI for p :
$0.933 \pm 1.96 \sqrt{\frac{0.933(1-0.933)}{104}}=(0.885,0.981)$
A traditional $95 \% \mathrm{CI}$ for p is $(0.907,0.993)$

